

Theorem 9 Define $y_n : \mathcal{J} \rightarrow D$ by the equality $y_n((-\infty, t)) = h_n(\Delta_t^n)$. Then y_n is an observable.

Proof. It follows by properties 1 - 3 of the previous Definition.

Definition 10 Let $x_1, \dots, x_n : \mathcal{J} \rightarrow D$ be independent observables. Then the observable $y_n : \mathcal{J} \rightarrow D$ defined in previous Theorem is called the sum of observables $x_1, \dots, x_n, y_n = \sum_{i=1}^n x_i$, i.e.

$$\left(\sum_{i=1}^n x_i\right)((-\infty, t)) = h_n(\Delta_t^n), t \in R.$$

Remark. There has been proved in [5] that in so-called Kôpka D-posets there exists the mapping $h_n : \mathcal{M}_n \rightarrow D$ satisfying the properties stated in previous Definition.

4. Central limit theorem

We are able now to formulate and prove the main result of the paper. We shall use the following notation. If $y : \mathcal{J} \rightarrow D$ is an observable and α, β are real numbers, $\alpha \neq 0$, then $\alpha y + \beta : \mathcal{J} \rightarrow D$ is defined by the formula

$$(\alpha y + \beta)((-\infty, t)) = y((-\infty, \frac{1}{\alpha} - \beta)).$$

Theorem 11 Let $(x_n)_{n=1}^\infty$ be an independent sequence of square integrable observables, $E(x_n) = a$, $\sigma^2(x_n) = \sigma^2$, ($n = 1, 2, \dots$). Then for any $t \in R$

$$\begin{aligned} \lim_{n \rightarrow \infty} m \left(\left(\frac{\sqrt{n}}{\sigma} \sum_{i=1}^n x_i - a \right) ((-\infty, t)) \right) &= \Phi(t) = \\ &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^t e^{-\frac{x^2}{2}} dx. \end{aligned}$$

Proof. Denote $P_n = \lambda_{F_1} \times \dots \times \lambda_{F_n} : \mathcal{B}(R^n) \rightarrow [0, 1]$. Then (P_n) presents a consistent system of probability measures:

$$P_n(A \times R) = P_{n-1}(A), A \in \mathcal{B}(R), n \in N$$

By the Kolmogorov consistence theorem there exists $P : \sigma(\mathcal{C}) \rightarrow [0, 1]$ where

$$\mathcal{C} = \{A \subset R^N; A = \pi_n^{-1}(B), B \in \mathcal{B}(R_n), n \in N\}$$

such that

$$P(\pi_n^{-1}(B)) = P_n(B) = \lambda_{F_1} \times \dots \times \lambda_{F_n}(B)$$

for any $B \in \mathcal{B}(R^n), n \in N$. Define $\xi_n : R^N \rightarrow R$ by the formula

$$\xi_n((u_i)_{i=1}^\infty) = u_n.$$

Therefore

$$\left(\sum_{i=1}^n x_i\right)((-\infty, t)) = h_n(\Delta_t^n) = \lambda_{F_1} \times \dots \times \lambda_{F_n}(\Delta_t^n) =$$

$$= P_n(\Delta_t^n) = \{\omega; \xi_1(\omega) + \dots + \xi_n(\omega) < t\}.$$

We obtained

$$\begin{aligned} m \left(\left(\frac{\sqrt{n}}{\sigma} \left(\sum_{i=1}^n x_i - a \right) \right) ((-\infty, t)) \right) &= \\ &= m \left(\left(\sum_{i=1}^n x_i \right) (-\infty, \frac{\sigma}{\sqrt{n}}(a+t)) \right) = \\ &= P \left(\{\omega; \frac{\sqrt{n}}{\sigma} \left(\sum_{i=1}^n \xi_i(\omega) - a \right) < t\} \right). \end{aligned}$$

Now by the classical limit theorem we obtain

$$\begin{aligned} \lim_{n \rightarrow \infty} m \left(\left(\frac{\sqrt{n}}{\sigma} \sum_{i=1}^n x_i - a \right) ((-\infty, t)) \right) &= \\ \lim_{n \rightarrow \infty} m \left(\left(\frac{\sqrt{n}}{\sigma} \sum_{i=1}^n \xi_i - a \right) ((-\infty, t)) \right) &= \Phi(t). \end{aligned}$$

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